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The Expected Number of Patterns in a Random Generated Permutation on \([n] = \{1,2,...,n\}\)

Evelyn Fokuoh

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The Expected Number of Patterns in a Random Generated Permutation on 

\[ [n] = \{1, 2, \ldots, n\} \]

A thesis

presented to

the faculty of the Department of Mathematics

East Tennessee State University

In partial fulfillment

of the requirements for the degree

Master of Science in Mathematical Sciences

by

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ABSTRACT

The Expected Number of Patterns in a Random Generated Permutation on 

\[ [n] = \{1, 2, \ldots, n\} \]

by

Evelyn Fokuoh

Previous work by Flaxman (2004) and Biers-Ariel et al. (2018) focused on the number of distinct words embedded in a string of words of length \( n \). In this thesis, we will extend this work to permutations, focusing on the maximum number of distinct permutations contained in a permutation on \([n] = \{1, 2, \ldots, n\}\) and on the expected number of distinct permutations contained in a random permutation on \([n]\). We further considered the problem where repetition of subsequences are as a result of the occurrence of (Type A and/or Type B) replications. Our method of enumerating the Type A replications causes double counting and as a result causes the count of the number of distinct sequences to go down.
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1 INTRODUCTION

A major problem of interest in combinatorics is the number of distinct subsequences that exist as subsequences of a given string. The number of distinct subsequence equals $2^n$ – the number of repetitions. For example, in the word 10110, the embedded subsequences are {}, 1, 0, 10, 01, 11, 101, 100, 110, 111, 1011, 1010, 1110, 0110, and 10110, where throughout this thesis we use the notation {} for empty set. There are thus 16 distinct subsequences among the $2^5 = 32$ subsequences of 10110. Similarly the permutation 1324 contains (using order-isomorphic representations we can, e.g., rename both 13 and 12 as 12) the sub-permutations {}, 1, 12, 21, 132, 123, 213, and 1324. The rest of the subpermutations are repetitions. Flaxman et al. (2004) focused on determining the maximizing string in the case of binary words. This turned out to be the alternating string. In other words, for the alphabet $\Sigma = \{1, 2\}$, the sequence $(1, 2, 1, 2, \ldots)$, gives the maximum number of distinct subsequences, which turn out to be

$$Fib(n + 3) - 1 \sim ((1 + \sqrt{5})/2)^{n+3}/\sqrt{5},$$

which is asymptotic to $(1.62)^n$. This result was generalized for a finite alphabet $\Sigma$ of size $d$. Biers-Ariel et al. (2018) defined $T_n$ as a fixed binary string of length $n$, $t_i$ as the $i^{th}$ letter of $T_n$, $S_n$ as a random binary string of length $n$ and $\phi(T_n)$ as the number of distinct subsequences of $T_n$. The authors of this paper focused on $E[\phi(S_n)]$ when $\Pr[s_i = 1] = \alpha \in [0, 1]$. The minimum value of $E(\phi(S_n))$ on a fixed length-$n$ string on the alphabet $\{0, 1\}$, is trivial. It occurs when $\alpha = 0$ or $\alpha = 1$, and in this case $E(\phi(S_n)) = n + 1$. If $n = 2, \alpha = 1/2$, the four possibilities for $T_n$
are 11, 10, 01, and 00. These contain 3, 4, 4, and 3 subsequences respectively, so 
\[ E(S_n) = \frac{3 + 4 + 4 + 3}{4} = 3.5. \] A key result in Biers-Ariel et al. (2018), which enables one to count \( \phi(T_n) \) recursively, is the following: Given \( T_n \), let \( l \) be the greatest number less than \( n \) such that \( t_l = t_n \), and if no such number exists, let \( l = 0 \). Then, \( \nu(T_n) \), the number of distinct new subsequences created at the \( n \)th stage in a string \( T_n \), is

\[
\nu(T_n) = \begin{cases} 
  n & \text{if } l = 0 \\
  \sum_{i=l}^{n-1} \nu(T_i) & \text{if } l > 0 
\end{cases}
\]

As an example, we consider the word \{10010\} of length \( n = 5 \). We have \( \phi(T_5) = 1 + \nu(T_1) + \nu(T_2) + \nu(T_3) + \nu(T_4) + \nu(T_5) = 1 + 1 + 2 + 2 + 5 + 7 = 18 \). Biers-Ariel et al. (2018) then used the above expression and linearity of expectation to find \( E(\phi(S_n)) \) for each \( \alpha \in (0,1) \). Later in the article, the authors focused on random strings from an alphabet of size \( d \), where the letter \( j \in [d] \) is independently selected with probability \( \alpha_j \). In other words, \( \Pr[s_i = j] = \alpha_j \) for all \( i \in [n], j \in [d] \). (Note that \( \sum_{j=1}^{d} \alpha_j = 1 \)). They gave a recursion to count \( \nu(T_n) \) as in the case of \( d = 2 \). They also considered the case where the letters were generated according to a Markov chain.

To give a specific example of a result in Biers-Ariel et al. (2018), suppose we are considering the case of binary words with unequal probabilities, i.e., \( \Pr[s_i = 1] = \alpha \in [0,1]; \Pr[s_i = 0] = 1 - \alpha \) for all \( 1 \leq i \leq n \). Then we have
\[ E[\phi(S_n)] = \begin{cases} 
 n + 1 & \text{if } \alpha = 0, 1 \\
 \frac{(1 - 2\sqrt{\alpha(1-\alpha)}) (1 - (1 - \sqrt{\alpha(1-\alpha)})^n) + (1 + 2\sqrt{\alpha(1-\alpha)}) ((1 + \sqrt{\alpha(1-\alpha)})^n - 1)}{2\sqrt{\alpha(1-\alpha)}} & \text{if } \alpha \neq 0, 1 
\end{cases} \]

Thus, \( E[\phi(S_n)] = 2\left(\frac{3}{2}\right)^n - 1 \) when \( \alpha = 0.5 \), as proved by Flaxman et al. (2004).

The main open problem stated in Biers Ariel et al. (2018) is extending the above work to permutations. Both the referees of their paper were enthusiastic about this direction.

In this thesis, \( \phi(\pi_n) \) is the number of distinct permutations contained in a permutation \( \pi_n \) on \([n]\). \( \pi_n \) may be fixed or random. In the random case, we can consider \( E(\phi(\pi_n)) \). \( \nu(j) \) is the number of new permutations created by the \( j \)th entry of the sequence \((\pi(1), \ldots, \pi(n))\) for \( j = (1, 2, \ldots, n) \). Finally, \( \psi(k, \pi) \) is the number of repeats of permutations of length \( k \). Some of the basic definitions of interest are the Type A repeats which are repeats in which two subpermutations have the same pattern and the values are identical except for two values which have the same rank in the pattern, e.g., subpermutations 2146 and 2145 of length \( k = 4 \). Both generate 2134 patterns. If \( n \) is odd, then the singleton \( \{n\} \) is a type C duplicate. Any duplicate which is neither a Type A nor a Type C is a type B duplicate.

Our two main directions will be the following:

(A) To determine the maximum number of embedded permutations as done by Flaxman et al. (2004) for words, since the minimum number of embedded permutations is trivial. (The permutation \( \pi_n = 123\ldots n \) contains the permutations \( \{\}, 1, 12, \ldots, 123\ldots n, \text{ or } n + 1 \text{ permutations} \).

(B) In the random case to make progress in determining \( E(\phi(\pi_n)) \).

These two topics will be presented in Sections 3 and 4 respectively, after some
preliminary results are presented in Section 2. Other issues such as the variance of \( \phi(\pi_n) \) or the distribution of \( \phi(\pi_n) \) are very difficult and we will not consider these.

Trying to find a recursion that counts the number of new permutations created at the \( j \)th position (this was the approach taken by Biers-Ariel et al. 2018) will be an option; this will hopefully give an exact expression for \( E(\phi(\pi_n)) \). Another option might be to show that \( E(\phi(\pi_n))/a^n \to 1 \) for some \( a \) (using subadditivity and Fekete’s Lemma). We will explore all these options in the coming pages.
2 PRELIMINARY RESULTS

In this part of the thesis we determine \((E[\phi(n)])^{1/n}\) for \(n = \{1, 2, 3, 4, 5, 6, 7, 8\}\). \(\nu(j)\) is defined to be the number of new permutations created by the \(j\)th entry of the sequence \((\pi(1), \ldots, \pi(n))\) for \(j = (1, 2, \ldots, n)\) and \(\phi(n)\) is the number of distinct permutations contained in the sequence \((\pi(1), \pi(2), \ldots, \pi(n))\). Hence \(\phi(n) = \sum_{j=1}^{n} \nu(j) + 1\). This was the approach taken by Biers-Ariel et al. (2018) for words, but after several attempts at finding analogs in the case of permutations, we abandoned this line of investigation. For small values of \(n\), we were able to use R (a free software for statistical computing) to write an R code. This is available in the Appendix. We derived histograms for \(n = \{1, 2, 3, 4, 5, 6, 7, 8\}\). The R on the x-axis defines the number of embedded sequences for each value of \(n\) while the y-axis defines the rate at which each embedded sequence occur (count). For instance in Figure 1, the possible permutation for \(n = 1\) is 1. The number of embedded subsequence when \(n = 1\) is \(\{\{\}, 1\}\). These two appear just once. For Figure 1, the mean of the embedded sequences when \(n = 1\) is 2. From our calculation, we had the variance of these sequences to be 0, the expected number of distinct permutations \((E[\phi(1)])^{1/1}\) was 2 and the value of the constant C was 1 (Mean/2\(^n\)). For Figure 2, with \(n = 2\), we can generate two possible permutations namely 12 and 21. The number of embedded sequences for 12 is \(\{\{\}, 1, 12\}\) and for 21 is \(\{\{\}, 2, 21\}\). However, by order isomorphism the 2 in 21 turns to be 1. Since we have three embedded sequences in both 12 and 21, it means the rate(frequency) at which we see the count of 3 is 2 times. These embedded sequences give us a mean value of 3, and a variance of 0.222. The \((E[\phi(2)])^{1/2}\) is 1.732 and the constant C is 0.75.
Now for Figure 3, where $n = 3$, the possible permutations are $123, 132, 213, 213, 231, 321$. There are 4 embedded sequences including the empty set in $123$ and $321$ but the rest contains 5 embedded sequences. Thus the rate at which we see the count of 4 is 2 and the rate at which we see the count of 5 is 4. The mean value of these embedded sequences when $n = 3$ is $4.667$, the variance is $0.222$, $(E[\phi(3)])^{1/3}$ is $1.671$ and the value of the constant $C$ is $0.5833$. This idea is used to explain the nature of the histograms, mean etc for Figures 4 through Figure 8. Detailed information on the embedded sequences for $n = \{1, 2, ..., 8\}$ can be found in the Rcode under Appendix.

Figure 1: A Histogram showing the value of $(E[\phi(n)])$ at different values of R when $n = 1$. 

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Figure 2: A Histogram showing the value of \( \mathbb{E}[\phi(n)] \) at different values of \( R \) when \( n = 2 \).

The mean for the embedded sequences for Figure 2 where \( n = 2 \) is 3 and the variance is 0. The \( \mathbb{E}[\phi(n)]^{1/n} \) and the value of the constant \( C \) are 1.732, 0.75 respectively.

Figure 3: A Histogram showing the value of \( \mathbb{E}[\phi(n)] \) at different values of \( R \) when \( n = 3 \).

For the Figure 3, the mean of the 6 embedded sequences is 4.667 and the variance is 0.222. The \( \mathbb{E}[\phi(n)]^{1/n} = 1.671 \) and \( C = 4.667/2^3 = 0.5833 \).
Figure 4: A Histogram showing the value of $(E[\phi(n)])$ at different values of R when $n = 4$.

When $n = 4$, the mean of the 24 embedded sequences is 7.417 and the variance = 0.9097. The $(E[\phi(n)])^{1/n} = 1.6502$ and $C = 7.417/2^4 = 0.4636$.

Figure 5: A Histogram showing the value of $(E[\phi(n)])$ at different values of R when $n = 5$.

For Figure 5, the mean of the 120 embedded sequences is 12.2833 and the variance = 4.1969. The $(E[\phi(n)])^{1/n} = 1.6514$ and $C = 12.283/2^5 = 0.3838$. 
Figure 6: A Histogram showing the value of $(\mathbb{E}[\phi(n)])$ at different values of R when $n = 6$.

The mean of the 720 embedded sequences in Fig. 6 is 21.1444 with a variance of 25.6347. $(\mathbb{E}[\phi(n)])^{1/n} = 1.6629$ and $C = \frac{21.1444}{2^6} = 0.3304$.

Figure 7: A Histogram showing the value of $(\mathbb{E}[\phi(n)])$ at different values of R when $n = 7$.

We generate 5040 embedded sequences when $n = 7$ with a mean of 37.60397 and variance of 76.0793. The $(\mathbb{E}[\phi(n)])^{1/n} = 1.6789$ and $C = \frac{37.604}{2^7} = 0.2938$. 
Figure 8: A Histogram showing the value of $\mathbb{E}[\phi(n)]$ at different values of $R$ when $n = 8$.

The mean for the embedded sequences is 68.63661 with a variance of 201.3242. The $\left(\mathbb{E}[\phi(n)]\right)^{1/n} = 1.6966$ and the constant $C = 68.637/2^8 = 0.2681$

We observe that the $\left(\mathbb{E}[\phi(n)]\right)$ was 2 for $n = 1$, then it decreased to 1.732 for $n = 2$. Then it reduced to 1.671 for $n = 3$. However after $n = 4$, it started to increase and finally to 1.6966 for $n = 8$. The value of the constant $C$ on the other hand was reducing for values of $n$ from 1 through 8. Our interest in deriving the values of the constant $C$ was because we believe $\mathbb{E}[\phi(n)] = C_n \cdot 2^n$. Hence, for reasons that we will give later, we believe that the following holds:

**Conjecture 2.1.** Since $\mathbb{E}[\phi(n)]$ is increasing after $n = 4$ and approaching a constant number for values of $n = \{1, 2, \ldots, 8\}$, then

$$\lim_{n \to \infty} \left(\mathbb{E}[\phi(n)]\right)^{1/n} = 2.$$
2.1 Possible use of Fekete’s Lemma (Subadditivity)

Suppose we can show that for each \( \pi \in S_{n+m} \), where \( S_n \) is the set of all permutations on \([n]\), we have that \( \phi(\pi_{m+n}) \leq \phi(\pi_n)\phi(\pi_{(n+1,...,n+m)}) \), then we would have made great progress. This was what occurred in Biers-Ariel et al (2018); they exploited subadditivity to prove the existence of a limit for \((E(\phi_n))^{1/n}\).

**Theorem 2.2.** Fekete’s Lemma: If \( a_n \) is a real sequence for which

\[
a_{n+m} \leq a_n + a_m, \quad (n, m = 1, 2, \ldots)
\]

then

\[
\lim_{n \to \infty} \frac{a_n}{n} = \inf \frac{a_n}{n},
\]

where the existence of the limit is the key conclusion.

As an example, we can consider \( \pi = 24315 \ n = 3 \) and \( m = 2 \). Then, \( \phi(\pi_{n+m}) = 14 \) due to the embedded subpermutations

\[
\{\}, 1, 12, 21, 123, 132, 231, 213, 321, 3214, 2314, 1324, 2431, 24315;
\]

\( \phi(\pi_n) = 5 \) due to

\[
\{\}, 1, 12, 21, 132;
\]

and \( \phi(\pi_{(n+1,...,n+m)}) = 3 \) due to

\[
\{\}, 1, 12.
\]

The inequality holds. If this was to be true for any \( \Pi \), then we could say that in general

\[
E[\phi_{n+m}] \leq E(\phi_n, \phi_{(n+1,...,n+m)}).
\]
Since this is on disjoint intervals, these values will be independent of each other. If we assume that \( n \leq m \) then we would have

\[
E(\phi_{n+m}) \leq E(\phi_n, \phi_{\{n+1, \ldots, n+m\}})
= E(\phi_n)E(\phi_{\{n+1, \ldots, n+m\}})
= E(\phi_n)E(\phi_m)
\]

(1)

Then,
\[
\log E(\phi_{n+m}) \leq \log E(\phi_n) + \log E(\phi_m)
\]

(2)

Letting \( a_n = \log E(\phi_n) \), this would lead to the conclusion that

\[
a_{n+m} \leq a_n + a_m,
\]

and thus, via Fekete’s Lemma, that

\[
\lim (E(\phi_n))^{1/n} = D.
\]

This is because we would have had

\[
\frac{a_n}{n} = \frac{1}{n} \log E(\phi_n) \to C,
\]

and thus

\[
\log E(\phi_n)^{\frac{1}{n}} \to C,
\]

which would imply that

\[
E(\phi_n)^{\frac{1}{n}} \to e^C := D.
\]

This was however a failed attempt. We actually found examples where the inequality didn’t hold. In an attempt to fix this, we examined the Erdős-deBruijn generalization of Fekete’s lemma.
2.2 Erdős-Debruijn Generalisation of Fekete’s Lemma

**Theorem 2.3.** Fekete’s Lemma: If we have \( \phi(n+m) \leq \phi(m) + \phi(n) + a(m+n), \)
where \( \sum_{n=1}^{\infty} \frac{a_n}{n^2} < \infty, \) then \( \lim_{n \to \infty} \frac{a_n}{n} \) exists.

We consider as examples the cases where \( \phi_{m+n} \leq \phi_n + \phi_m + n + m, \) and \( \phi_{m+n} \leq \phi_n + \phi_m + \sqrt{n+m}. \) Theorem 3 could not be used in the first case but could be used in the second case, since the series \( \sum \frac{1}{n} \) diverges but the series \( \sum \frac{1}{n^{3/2}} \) converges. Our efforts to find a “buffer” function that led to Theorem 3 being valid were unsuccessful as well.

2.3 Trying to use a Recursion Similar to that in Biers-Ariel et al (2018)

We have that

\[
\phi(\pi_n) = \sum_{j=1}^{n} \nu(j)
\]

where \( \nu(j) \) is the number of new permutations counted by the \( j \)th entry. We can calculate \( \nu(j) \) as follows: If \( \nu(j) \) is the smallest term, increase all the previous entries by 1 and label the \( j \)th entry as 1. Retain all those sequences that have not occurred before. If we consider as an example the permutation 42681, the possible
subpermutations are:

\[
\begin{align*}
  j & = 1; \nu(j) = 1; 1 \\
  j & = 2; \nu(j) = 1; 21 \\
  j & = 3; \nu(j) = 2; 12, 213 \\
  j & = 4; \nu(j) = 2; 23, 2134 \\
  j & = 5; \nu(j) = 5; 321, 231, 3241, 32451, 2341.
\end{align*}
\]

(3)

We can observe that the last entry \( j = 5 \) was obtained by increasing all previous entries we have obtained by 1 and finally ending with a term 1.

If \( \nu(j) \) is the largest term encountered, we augment each previous term by the next largest value and retain all those sequences that have not occurred before. We take a look at the permutation 42689:

\[
\begin{align*}
  j & = 1; \nu(j) = 1; 1 \\
  j & = 2; \nu(j) = 1; 21 \\
  j & = 3; \nu(j) = 2; 12, 213 \\
  j & = 4; \nu(j) = 2; 123, 2134 \\
  j & = 5; \nu(j) = 2; 1234, 21345.
\end{align*}
\]

(4)

However, this process is very difficult to track when terms other than the largest or smallest appear at the \( j \)th position. We thus abandoned our efforts to adopt a
method similar to that in Biers-Ariel et al. (2018), which was used so successfully for words.

2.4 Heuristic Reason why we think $E(\phi_n)^{1/n} \rightarrow 2$

We see from the histograms that $E(\phi_n)^{1/n}$ is decreasing up till $n = 5$ and increasing after that, up till $n = 8$. Will it approach 2 as $n \to \infty$? Here we argue that it might be possible, and the rest of the thesis is devoted to gathering evidence that the limit might equal 2.

Now for large $k$, we observed that $k! \gg \binom{n}{k}$. To obtain the repeats of length $k$ for large $k$, we are essentially asking how many repeats are possible in a small set of $\binom{n}{k}$ positions, where there exist $k!$ possible permutations. Intuition suggests that this would be a small number. This thesis will be devoted to classifying and understanding these repeats.

Let the number of repeats of permutations of length $k$ be denoted by $\psi(k, \pi)$. Suppose that for a large $k$, say $k \geq n/2$, we have at most $\binom{n}{k}/2$ repeats, so that $\max \psi(k, \pi) \leq \binom{n}{k}/2$. Then the total number of repeats is

$$
\sum_{k=1}^{n/2} \max \psi(k, \pi) + \sum_{n/2+1}^{n} \max \psi(k, \pi)
$$

which is bounded above by

$$
\sum_{k=1}^{n/2} \binom{n}{k} + \sum_{n/2+1}^{n} \binom{n}{k}/2,
$$

which simplifies to

$$
\frac{1}{2} \times 2^n + \frac{1}{2} \times 2^n \times \frac{1}{2} = \frac{3}{4} \times 2^n.
$$
Therefore, the number of distinct permutations would be at least $2^n - \frac{3}{4} \times 2^n = \frac{1}{4} \times 2^n$, and hence we would have that

$$E(\phi_n)^{1/n} \geq \frac{1}{4^{1/n}} \times 2 \rightarrow 2.$$ 

2.5 Types of Repeats

In the permutation $n = 2146537$, let $k = 4$. Then we have $k! = 24$ and $\binom{n}{k} = 35$. The successive terms are (1, 2) and (5, 6). The possible subpermutations of length 4 from the above permutation $n = 2146537$ are obtained in the Table 1 below. From Table 1, the colored pattern under the reduced column indicate some of the Type A repeat. In this table, if we consider case 1 and case 2, thus the subpermutations 2146 and 2145. These two subpermutations generate the same reduce pattern 2134. This is because 6 in the subpermutation 2146 and 5 in subpermutation 2145 turn to have the same rank (which is 4) in the reduced pattern. Again from case 4 and case 7, the subpermutations 2147 and 2167 both generate a reduce pattern of 2134. This is because 4 in the subpermutation 2147 and 6 in the subpermutation 2167 have the same rank of 3 in the reduced pattern. The rest of the Type A repeats are found in the Table 1 below and those that can be paired are colored with the same color.
Table 1: Types of repeats.

<table>
<thead>
<tr>
<th>CASE</th>
<th>SUBPERMUTATION</th>
<th>REDUCTION</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>2146</td>
<td>2134</td>
</tr>
<tr>
<td>2</td>
<td>2145</td>
<td>2134</td>
</tr>
<tr>
<td>3</td>
<td>2143</td>
<td>2143</td>
</tr>
<tr>
<td>4</td>
<td>2147</td>
<td>2134</td>
</tr>
<tr>
<td>5</td>
<td>2165</td>
<td>2143</td>
</tr>
<tr>
<td>6</td>
<td>2163</td>
<td>2143</td>
</tr>
<tr>
<td>7</td>
<td>2167</td>
<td>2134</td>
</tr>
<tr>
<td>8</td>
<td>2153</td>
<td>2143</td>
</tr>
<tr>
<td>9</td>
<td>2157</td>
<td>2134</td>
</tr>
<tr>
<td>10</td>
<td>2137</td>
<td>2134</td>
</tr>
<tr>
<td>11</td>
<td>2465</td>
<td>1243</td>
</tr>
<tr>
<td>12</td>
<td>2463</td>
<td>1342</td>
</tr>
<tr>
<td>13</td>
<td>2467</td>
<td>1234</td>
</tr>
<tr>
<td>14</td>
<td>2453</td>
<td>1342</td>
</tr>
<tr>
<td>15</td>
<td>2457</td>
<td>1234</td>
</tr>
<tr>
<td>16</td>
<td>2437</td>
<td>1324</td>
</tr>
<tr>
<td>17</td>
<td>2653</td>
<td>1432</td>
</tr>
<tr>
<td>18</td>
<td>2657</td>
<td>1324</td>
</tr>
<tr>
<td>19</td>
<td>2637</td>
<td>1324</td>
</tr>
<tr>
<td>20</td>
<td>2537</td>
<td>1324</td>
</tr>
<tr>
<td>21</td>
<td>1465</td>
<td>1243</td>
</tr>
<tr>
<td>22</td>
<td>1463</td>
<td>1342</td>
</tr>
<tr>
<td>23</td>
<td>1465</td>
<td>1234</td>
</tr>
<tr>
<td>24</td>
<td>1453</td>
<td>1342</td>
</tr>
<tr>
<td>25</td>
<td>1457</td>
<td>1234</td>
</tr>
<tr>
<td>26</td>
<td>1437</td>
<td>1324</td>
</tr>
<tr>
<td>27</td>
<td>1653</td>
<td>1432</td>
</tr>
<tr>
<td>28</td>
<td>1657</td>
<td>1324</td>
</tr>
<tr>
<td>29</td>
<td>1637</td>
<td>1324</td>
</tr>
<tr>
<td>30</td>
<td>1537</td>
<td>1324</td>
</tr>
<tr>
<td>31</td>
<td>4653</td>
<td>2431</td>
</tr>
<tr>
<td>32</td>
<td>4657</td>
<td>1324</td>
</tr>
<tr>
<td>33</td>
<td>4637</td>
<td>2314</td>
</tr>
<tr>
<td>34</td>
<td>4537</td>
<td>2314</td>
</tr>
<tr>
<td>35</td>
<td>6537</td>
<td>3214</td>
</tr>
</tbody>
</table>
The only type $C$ present in this permutation is 7. We start by analyzing type $A$ repeats with successions. A succession is defined as an occurrence where $\pi(i + 1) = \pi(i) \pm 1$.

There are however other types of events that cause duplication other than successions. In this thesis, we focus on what we call Type A duplicates, which are defined more precisely (than done before) later.

2.6 Calculation of the Expected Number of Patterns in a Random Permutation

We start by calculating some values of $[E(\phi(n))]^{1/n}$ by hand and by using the R code found in the Appendix. The expected number of patterns for $n = 2, 3, 4$ are in the tables below:

From the Table 2, we derive two possible permutations of length 2. This is found in the first column. The number of patterns is 3 as explained for Figure 2 for each of these permutations. The probability that each of these permutations occur is 0.5. The $[E(\phi(2))]^{1/2} = (6/2)^{0.5} = 1.7321$. Hence, the expected number of patterns for $n = 2$ is 1.7321.

<table>
<thead>
<tr>
<th>$\pi_n$</th>
<th>number of patterns ($\phi_n$)</th>
<th>Probability</th>
</tr>
</thead>
<tbody>
<tr>
<td>12</td>
<td>{ },1,12 =3</td>
<td>1/2=0.5</td>
</tr>
<tr>
<td>21</td>
<td>{ },1,21 =3</td>
<td>0.5</td>
</tr>
</tbody>
</table>
For $n = 3$, the column 1 of Table 3 gives us information on all the 6 possible permutations for $n = 3$. This is similar to the number of permutations we derived for Figure 3. The $[E(\phi(3))]^{1/3} = (28/6)^{1/3} = 1.671$. Since we have 6 permutations in all, the probability that each of these permutations occur is 0.167.

Table 3: Expected number of patterns for $n=3$.

<table>
<thead>
<tr>
<th>$\pi_n$</th>
<th>number of patterns ($\phi_n$)</th>
<th>Probability</th>
</tr>
</thead>
<tbody>
<tr>
<td>123</td>
<td>${},1,12,123=4$</td>
<td>1/6=0.167</td>
</tr>
<tr>
<td>132</td>
<td>${},1,12,21,132=5$</td>
<td>0.167</td>
</tr>
<tr>
<td>213</td>
<td>${},1,21,12,213=5$</td>
<td>0.167</td>
</tr>
<tr>
<td>231</td>
<td>${},1,12,21,231=5$</td>
<td>0.167</td>
</tr>
<tr>
<td>312</td>
<td>${},1,21,12,312=5$</td>
<td>0.167</td>
</tr>
<tr>
<td>321</td>
<td>${},1,21,321=4$</td>
<td>0.167</td>
</tr>
</tbody>
</table>

For $n = 4$, we have a total of 24 possible permutations of length 4 each in the Table 4. $[E(\phi(4))]^{1/4} = (178/24)^{1/4} = 1.6502$.

Table 4: Expected number of patterns for $n=4$.

<table>
<thead>
<tr>
<th>$\pi_n$</th>
<th>number of patterns ($\phi_n$)</th>
<th>Contribution to Expected Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>1234</td>
<td>${},1,12,123,1234=5$</td>
<td>5*1/24=5/24</td>
</tr>
<tr>
<td>1243</td>
<td>${},1,12,21,123,132,1243=7$</td>
<td>7/24</td>
</tr>
<tr>
<td>1324</td>
<td>${},1,12,21,123,132,213,1324=8$</td>
<td>8/24</td>
</tr>
<tr>
<td>1342</td>
<td>${},1,12,21,123,132,231,1342=8$</td>
<td>8/24</td>
</tr>
<tr>
<td>1423</td>
<td>${},1,12,21,312,132,123,1423=8$</td>
<td>8/24</td>
</tr>
<tr>
<td>1432</td>
<td>${},1,12,21,321,1432=7$</td>
<td>7/24</td>
</tr>
<tr>
<td></td>
<td>$\vdots$</td>
<td>$\vdots$</td>
</tr>
<tr>
<td>4312</td>
<td>${},1,12,21,312,321,4312=7$</td>
<td>7/24</td>
</tr>
<tr>
<td>4321</td>
<td>${},1,21,321,4321=5$</td>
<td>5/24</td>
</tr>
</tbody>
</table>
The probability that each of these permutations in Table 4 will occur is $1/24$. However, the contribution of $5$ to the expected value is $5/24$.

The corresponding values for $[E(\phi(n))]^{1/n}$ for $n = 5, 6$ and $7$ are:

$[E(\phi(5))]^{1/5} = (1, 474/120)^{1/5} = 1.6514$.

$[E(\phi(6))]^{1/6} = (15, 224/720)^{1/6} = 1.66289$.

$[E(\phi(7))]^{1/7} = (189, 524/5040)^{1/7} = 1.6789$.

We see that the minimum of $[E(\phi(n))]^{1/n}$ is at $n = 4$, and we observe an increasing trend from there on.
3 UPPER AND LOWER BOUNDS ON $\phi(n)$

It is clear that there are at most $k!$ distinct permutations of length $k$. But only $\binom{n}{k}$ of these may actually be present. For small values of $k$, $k! < \binom{n}{k}$, and this inequality flips for large values of $k$. As an example we consider $n = 7$. The values of $k!$ are $(1, 2, 6, 24, 120, 720, 5040)$, and the corresponding values for $\binom{n}{k}$ are $(7, 21, 35, 35, 21, 7, 1)$. In any case, we have the bound

$$\phi(n) \leq \sum_{k=0}^{n} \min\left(\binom{n}{k}, k!\right).$$

We believe (and will argue below) that $\sum_{k=0}^{n} \min\left(\binom{n}{k}, k!\right)$ is not a bad bound. To explain this we take a look at $\sum_{k=0}^{n} \min\left(\binom{n}{k}, k!\right)$. We start by analyzing the relationship between $\binom{n}{k}$ and $k!$:

By Stirling’s approximation

$$k! \sim \sqrt{2\pi k} \left(\frac{k}{e}\right)^k,$$

so on taking the limit we obtain

$$\lim_{k \to \infty} \frac{k!}{\sqrt{2\pi k} \left(\frac{k}{e}\right)^k} = 1.$$

For any $k$, however, we have the bounds

$$\sqrt{2\pi k} \left(\frac{k}{e}\right)^k \leq k! \leq \sqrt{2\pi k} \left(\frac{k}{e}\right)^k e^{\frac{1}{12k}}.$$

The other term is $\binom{n}{k}$. If $k$ is not of magnitude $O(n)$, (and thus $k \neq \frac{n}{3}$, for example), then we can approximate

$$\frac{(n-k)^k}{k!} \leq \binom{n}{k} \leq \frac{n^k}{k!}.$$
and thus
\[
\binom{n}{k} \sim \frac{n^k}{\sqrt{2\pi k \left( \frac{k}{e} \right)^k}}.
\]

Since
\[
k! \sim \sqrt{2\pi k} \left( \frac{k}{e} \right)^k,
\]
on equating the two we obtain
\[
\sqrt{2\pi k} \left( \frac{k}{e} \right)^k = \frac{n^k}{\sqrt{2\pi k \left( \frac{k}{e} \right)^k}},
\]
or
\[
2\pi k \left( \frac{k^2}{ne^2} \right)^k = 1.
\]
In order to get a close estimate for the solution, we ignore the linear factor $2\pi k$, and obtain
\[
k^2 = \frac{ne^2}{1},
\]
or
\[
k = e\sqrt{n}.
\]
Albert et al (3) noted the bound $\sum \min (\binom{n}{k}, k!)$, but dismissed it without further analysis. Let us summarize the history of the lower bound on the maximum number of permutations $\max \phi(n)$ contained in an $n$ permutation:

Coleman (6) proved that
\[
\max \phi(n) \geq 2^{n-2\sqrt{n}+1}
\]
establishing that
\[
\left( \max \phi(n) \right)^{1/n} \to 2
\]
as \( n \to \infty \), but leaving open the question of whether

\[
\frac{\max \phi(n)}{2^n} \to 1.
\]

Albert’s et al’s lower bound on the maximum number of embedded distinct permutations was larger than what Coleman had. Specifically, they proved that

\[
\max \phi(n) \geq 2^n \left(1 - \frac{6\sqrt{n}}{2^{\frac{3}{2}}}ight) = 2^n (1 - \epsilon),
\]  
(5)

thus showing that

\[
\frac{\max \phi(n)}{2^n} \to 1.
\]

Miller (2008) proved a better lower bound than in (5). Specifically, she proved that

\[
\max \phi(n) \geq 2^n - O(n^{2n-\sqrt{2n}}),
\]

which she matched with the upper bound of

\[
\max \phi(n) \leq 2^n - \Theta(n^{2n-\sqrt{2n}}).
\]

(7)

We delved through the paper to uncover the value of the constant in (7) and found that it was enormous. With \( l = \sqrt{2n} \), (7) can be fleshed out as

\[
\max \phi(n) \leq 2^n - \left(2^n \left(\frac{nl - 27l - \frac{27n}{2} - 27}{80(l + 3)2^l}\right) - (2^{16}\cdot n^4\cdot 2^{(n-l)})\right).
\]

(8)

In the Table 5 below, we compare our upper bound \( \sum \min\left(\binom{n}{k}, k!\right) \) to Miller’s upper bound. We observed clearly the magnitude of the constant in (8) leads to the “trivial” bound, doing much better for values of \( n \) at least till 50.
Table 5: Comparison of our upper bound with Miller’s upper bound.

<table>
<thead>
<tr>
<th>$n$</th>
<th>Our upper bound</th>
<th>Miller’s upper bound</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>2</td>
<td>49,182.218</td>
</tr>
<tr>
<td>2</td>
<td>3</td>
<td>1,048,580.26</td>
</tr>
<tr>
<td>3</td>
<td>5</td>
<td>2,654,216.42</td>
</tr>
<tr>
<td>4</td>
<td>9</td>
<td>37,791,894.67</td>
</tr>
<tr>
<td>5</td>
<td>16</td>
<td>146,409,613.6</td>
</tr>
<tr>
<td>10</td>
<td>540</td>
<td>$3.02059167 \times 10^{10}$</td>
</tr>
<tr>
<td>100</td>
<td>$1267650591324628\ldots$</td>
<td>$4.594862602\times 10^{10}$</td>
</tr>
<tr>
<td>200</td>
<td>$160693804425899027\ldots$</td>
<td>$1.60693806\times 10^{28}$</td>
</tr>
</tbody>
</table>

Next, we study the asymptotics of the $\sum \min\left(\binom{n}{k}, k!\right)$ bound, and will uncover that while not as good as (8), does satisfy several desirable criteria. We have already seen that $k! = \binom{n}{k}$ around $k = e\sqrt{n}(1 + o(1)) = A$. We thus truncate as follows:

$$
\sum \min\left(\binom{n}{k}, k!\right) = \sum_{0}^{A} k! + \sum_{A+1}^{n} \binom{n}{k} \\
= \sum_{0}^{A} k! + \sum_{0}^{n} \binom{n}{k} - \sum_{0}^{A} \binom{n}{k} \\
= \sum_{0}^{n} \binom{n}{k} - \sum_{0}^{A} \left(\binom{n}{k} - k!\right) \\
= 2^n - \sum_{0}^{A} \left(\binom{n}{k} - k!\right) \\
= 2^n - \sum_{0}^{A} \binom{n}{k} + \sum_{0}^{A} k!.
$$

(9)
We need to find a good upper bounds on the two sums in (9). First we have

\[
\sum_0^A \binom{n}{k} = 2^n \sum_0^A \binom{n}{k} \frac{1}{2^k 2^{n-k}} \\
= 2^n \Pr(B_i(n, \frac{1}{2}) \leq A) \\
\geq 2^n \Pr(B_i(n, \frac{1}{2}) = A)(1 + \epsilon) \\
= \left(\frac{n}{A}\right)(1 + \epsilon),
\]

where we used, in (11), Proposition A.2.5 in Barbour et al (1992). Since \(A \sim e\sqrt{n}\), we see that

\[
\left(\frac{n}{A}\right) \sim \frac{n^A}{(\sqrt{n}e)!} \sim (\sqrt{n}e^2)^{\sqrt{n}e}.
\]

Plugging this into (11), we see that

\[
\sum \min\left(\binom{n}{k}, k!\right) \geq 2^n - (\sqrt{n}e^2)^{\sqrt{n}e} + \sum_0^A k! \\
\sim 2^n - (\sqrt{n}e^2)^{\sqrt{n}e} + A \cdot A! \\
= 2^n - 2^{c\sqrt{n}\ln n} + A \cdot A! \\
\sim 2^n - 2^{c\sqrt{n}\ln n} + C\sqrt{n} \cdot (\sqrt{n})^{e\sqrt{n}} \\
= 2^n - 2^{c\sqrt{n}\ln n}(1 + o(1)).
\]

Thus it turns out that asymptotically, Miller gets a better upper bound which is

\(2^n - (n2^{n-\sqrt{2n}})\). This is because

\[
2^n - (n2^{n-\sqrt{2n}}) \leq 2^n - 2^{c\sqrt{n}\ln n}
\]

since \(2^{c\sqrt{n}\ln n} \leq n2^{n-\sqrt{2n}}\). As seen however, our bound does far better for small values of \(n\).
4 ANALYSIS

For any \( n, k \), most of the duplicates are created by choosing one entry in a pair that has the same rank (all other terms are the same). We call such duplicates as Type A duplicates.

4.1 Interesting Examples

If we consider as an example the permutation 132, the possible patterns are in the Table 6 below,

<table>
<thead>
<tr>
<th></th>
<th>1</th>
<th>12</th>
<th>21</th>
<th>132</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>12</td>
<td>32</td>
<td>132</td>
<td></td>
</tr>
<tr>
<td>2</td>
<td>13</td>
<td>null</td>
<td>null</td>
<td></td>
</tr>
<tr>
<td>3</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td></td>
</tr>
</tbody>
</table>

Here the pair 1, 2 creates a duplicate since both have a rank of 1. In the second column, a duplicate is also generated from the pair 12 and 13. This is because the 2 and 3 in these terms have the same rank. When we pair these up the duplicate created is due to \((2, 3)\). This is true for all permutations on \( n = 2 \) from which we derive 2 possible permutation 12 and 21 in Tables 7 and 8 respectively. The Tables 9 through 14 gives us all possible patterns of the permutations on \( n = 3 \). Since \( n \) is odd, all the 3rd rows of the 1st column gives us type C repeats since they cannot be paired.
Table 7: Patterns in the permutation 12.

<table>
<thead>
<tr>
<th></th>
<th>1</th>
<th>12</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>1</td>
<td>12</td>
</tr>
<tr>
<td>2</td>
<td>null</td>
<td></td>
</tr>
</tbody>
</table>

Table 8: Patterns in the permutation 21.

<table>
<thead>
<tr>
<th></th>
<th>1</th>
<th>21</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>21</td>
<td>1</td>
</tr>
<tr>
<td>2</td>
<td>null</td>
<td></td>
</tr>
</tbody>
</table>

Table 9: Patterns in the permutation 123.

<table>
<thead>
<tr>
<th></th>
<th>1</th>
<th>12</th>
<th>123</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>12</td>
<td>1</td>
<td>123</td>
</tr>
<tr>
<td>2</td>
<td>13</td>
<td>null</td>
<td></td>
</tr>
<tr>
<td>3</td>
<td>23</td>
<td>null</td>
<td></td>
</tr>
</tbody>
</table>

Table 10: Patterns in the permutation 132.

<table>
<thead>
<tr>
<th></th>
<th>1</th>
<th>12</th>
<th>21</th>
<th>132</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>13</td>
<td>32</td>
<td>132</td>
<td></td>
</tr>
<tr>
<td>2</td>
<td>12</td>
<td>null</td>
<td>null</td>
<td></td>
</tr>
<tr>
<td>3</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td></td>
</tr>
</tbody>
</table>

Table 11: Patterns in the permutation 213.

<table>
<thead>
<tr>
<th></th>
<th>1</th>
<th>12</th>
<th>21</th>
<th>213</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>13</td>
<td>21</td>
<td>213</td>
<td></td>
</tr>
<tr>
<td>2</td>
<td>23</td>
<td>null</td>
<td>null</td>
<td></td>
</tr>
<tr>
<td>3</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td></td>
</tr>
</tbody>
</table>
Table 12: Patterns in the permutation 231.

<table>
<thead>
<tr>
<th></th>
<th>12</th>
<th>21</th>
<th>231</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>23</td>
<td>21</td>
<td>231</td>
</tr>
<tr>
<td>2</td>
<td>null</td>
<td>31</td>
<td>null</td>
</tr>
<tr>
<td>3</td>
<td>null</td>
<td>null</td>
<td>null</td>
</tr>
</tbody>
</table>

Table 13: Patterns in the permutation 312.

<table>
<thead>
<tr>
<th></th>
<th>12</th>
<th>21</th>
<th>312</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>12</td>
<td>31</td>
<td>312</td>
</tr>
<tr>
<td>2</td>
<td>null</td>
<td>32</td>
<td>null</td>
</tr>
<tr>
<td>3</td>
<td>null</td>
<td>null</td>
<td>null</td>
</tr>
</tbody>
</table>

Table 14: Patterns in the permutation 321.

<table>
<thead>
<tr>
<th></th>
<th>21</th>
<th>321</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>21</td>
<td>132</td>
</tr>
<tr>
<td>2</td>
<td>31</td>
<td>null</td>
</tr>
<tr>
<td>3</td>
<td>32</td>
<td>null</td>
</tr>
</tbody>
</table>

We now extend this to a larger $n$, using say 4132.

1. The old subsequence of size 3, before the 4 was added, was 132. In this term, the only length 3 permutation is 132.

2. The subsequences of size 2 from 132 are 13, 12, 32. After adding 4, we get 413, 412, 432 as additional pattern of length 3.
Table 15: Patterns in the permutation 4132.

<table>
<thead>
<tr>
<th></th>
<th>12</th>
<th>21</th>
<th>312</th>
<th>321</th>
<th>132</th>
<th>4132</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1</td>
<td></td>
<td></td>
<td>41</td>
<td>413</td>
<td>432</td>
<td>4132</td>
</tr>
<tr>
<td>2</td>
<td>12</td>
<td>43</td>
<td></td>
<td></td>
<td>null</td>
<td></td>
</tr>
<tr>
<td>3</td>
<td>null</td>
<td>42</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td></td>
</tr>
<tr>
<td>4</td>
<td>null</td>
<td>32</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td></td>
</tr>
</tbody>
</table>

From the Table 15 above, it can be observed that subsequences of size 1 become subsequences of size 2 when 4 is added; and subsequences of size 2 become subsequences of size 3. Hence the original pair of length 2 in the 2nd column turns to be a pair of length 3 in the 4th column. We can observe that subsequences of size 3 are derived from old subsequences of size 3 and old subsequences of size 2 plus the element $n$. Hence in general, the subsequences of size $k$ is obtained from old subsequences of size $k$ plus old subsequences of size $k - 1 + a$ new term, namely $n$. Thus by Pascal’s identity, we have

$$\binom{n}{k} = \binom{n - 1}{k} + \binom{n - 1}{k - 1}.$$  

If we consider the permutation $n = 41325$, the subsequences of size 3 we generate are 413, 412, 432, 132, 135, 125, 415, 435, 425, 325. The old subsequences of size 3 generated from 4132 are 413, 412, 432, 132. Again, the old subsequences of size 2 that we generate from 4132 are 13, 12, 41, 43, 42, 32. When we add the new term 5, each of these length of 2 permutations changes to 135, 125, 415, 435, 425 and 325. Hence the subsequences of length 3 from 41325 is the sum of the old subsequences of length 3 generated from 4132 and the old subsequences of length 2 from 4132 with 5 added to each pattern. This holds for larger values of $n$. 37
We have defined type A duplicates as those that occur when two terms of a permutation can be paired because they have the same rank in a permutation of length $k$ and thus yield the same pattern. Type B duplicates are those that remain after all the type A and type C duplicates have been enumerated. (Type C duplicates may occur as a result of the singletons for odd values of $n$.) We take a look at an example:

Table 16: Patterns in the permutation 41352.

<table>
<thead>
<tr>
<th></th>
<th>1</th>
<th>12</th>
<th>21</th>
<th>123</th>
<th>132</th>
<th>213</th>
<th>231</th>
<th>312</th>
<th>321</th>
<th>3124</th>
<th>4132</th>
<th>3142</th>
<th>3241</th>
<th>1342</th>
<th>41352</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>13</td>
<td>41</td>
<td>135</td>
<td>132</td>
<td>415</td>
<td>452</td>
<td>413</td>
<td>432</td>
<td>4135</td>
<td>4132</td>
<td>4152</td>
<td>4352</td>
<td>1352</td>
<td>41352</td>
<td></td>
</tr>
<tr>
<td>2</td>
<td>15</td>
<td>43</td>
<td>null</td>
<td>152</td>
<td>435</td>
<td>352</td>
<td>412</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td></td>
</tr>
<tr>
<td>3</td>
<td>12</td>
<td>42</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td></td>
</tr>
<tr>
<td>4</td>
<td>45</td>
<td>32</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td></td>
</tr>
<tr>
<td>5</td>
<td>35</td>
<td>52</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td>null</td>
<td></td>
</tr>
</tbody>
</table>

In the first column of the above table (Table 16), we have duplicates generated by (1,2) and (3,4). All these form a rank of 1 and hence they can be paired. In the second column, we generate a duplicate of generated by the pair (3, 5) from (13, 15). This is a duplicate of type A since 3 and 5 have a same rank in these two permutations. The total number of permutations generated including {} is $2^n = 32$. After removing the pairings we have $32 - 2(10) = 12$. Here we have 10 paired patterns that form a Type A and 12 remaining distinct permutations. The 5 in the first column is neither a Type A nor a Type B so it forms a Type C duplicate. We have 15 distinct permutations, but on removing the duplicates we see that we destroy some permutations, leading to only 12 remaining distinct permutations. From the above permutation, 5 is larger.
than all the other terms so the same pairing remains. Old subsets of size 2 are created by 12 and 13. New subsets of size 2 are created by adding 5 to all the old subsets of size 1. Old subsets of size 3 are 132, 432, 412, 413. The new subsets generated are 415, 425, 435, 135, 125, 325. By pairing, we realize that 425 and 435 can be paired because 2 and 3 are of the same rank. 413 and 412 are also pairs with 2 and 3 having the same rank. 135 and 125 can also be paired. We realize the permutation 432 cannot be paired.

Now the role played by Type B repeats is best seen in the permutation 456123. There are exactly two 123 patterns and they have no elements in common. Here is another example:

The permutation 312645 contain the subsets 315, 325, 164, 165, 364, 365. By pairing 315 and 325 we generate a 213 pattern because 1 and 2 turn to have the same rank. The permutation 164 and 165 also generate 132 pattern. 164 and 165 are pairs because 4 and 5 have the same rank. Here again, there are two 312 patterns namely 312 and 645 which do not have any elements in common.

We clearly need to better understand Type B duplicates. All repeats are either a type A or type B, except for the trivial case of up to one Type C duplicate. We therefore have \( E(\phi(n)) = 2^n - E(\psi(n)) \), where \( E(\phi(n)) \) and \( E(\psi(n)) \) are the expected number of distinct patterns and duplicates respectively.

Since all repeats are of type A or type B, and the singletons can be considered as a type C, our goal is to get an upper bound on \( E(\psi(n)) \). Where
\[
E(\psi(n)) \leq E(\psi_A(n)) + E(\psi_B(n)) + 1
\]

\[
= \sum_{b,c} E(\psi_{A,b,c}(n)) + \sum E(\psi_B(n)) + 1
\]

\[
= \sum_k \sum_{b,c} E(\psi_{A,b,c}(n,k)) + \sum E(\psi_B(n)) + 1
\]

\[
= \sum_1 + \sum_2 + 1. \quad (13)
\]

Here we let,

1. \(E(\psi_A(n))\) is the number of repeats caused by Type A and \(E(\psi_B(n))\) is the number of repeats caused by Type B.

2. \(\sum_{b,c} E(\psi_{A,b,c}(n))\) is the number of repeats of type A caused by \(b\) and \(c\) where \(b\) and \(c\) are the two different fixed points and \(\sum E(\psi_B(n))\) is the number of repeats caused by Type B situations.

3. \(\sum_k \sum_{b,c} E(\psi_{A,b,c}(n,k))\) is the number of repeats caused by \(b\) and \(c\) of length \(k\).

4. \(\sum_1\) denotes the duplicates of type A and \(\sum_2\) denotes the duplicates of type B.

An upper bound on \(\sum_1\) is given in the following main theorem.

**Theorem 4.1.**

\[
E \left( \sum_1 \right) = \sum_{k=1}^{n} \sum_{t=1}^{n-1} \sum_{r=0}^{n-2} \sum_{s=0}^{\min(t-1,r)} \frac{2(n-t)(n-r-1)}{n!} \times \left( \begin{array}{c} r \\ s \end{array} \right) \times \left[ (t-1) \ldots (t-1-(s-1)) \right] \times \left[ (n-r-2)(n-r-2-1) \ldots (n-r-2-(t-2-s)) \right] \times (n-2-t+1)! \left( \begin{array}{c} n-r-t-1+s \\ k-1 \end{array} \right) \quad (14)
\]
where

1. \( k \) is the number of repeats of type A of size \( k \) with the least being 1 and the maximum as \( n \).

2. \( t \) is the numerical difference between \( b \) and \( c \) and the least difference we can have is 1.

3. \( r \) is the position difference between \( b \) and \( c \) and the least is 0.

4. \( s \) the number of terms that are in value between \( b \) and \( c \) and are physically between \( b \) and \( c \).

5. 2 is the way we can permute the two numbers (\( b \) and \( c \) or \( c \) and \( b \)).

6. \((n - t)\) is the number of ways we can have the numerical difference being \( t \).

7. \((n - r - 1)\) is the number of ways we can place the two numbers.

8. \( n! \) defines the total number of permutation and it is needed since we are dealing with expected value.

9. \( \binom{r}{s} \) defines the positions of the \( s \) numbers, out of the \( r \) numbers, that are in value between \( b \) and \( c \).

10. \([ (t - 1) \ldots (t - 1 - (s - 1)) ] \) is the number of ways we can choose the \( s \) numbers.

11. \([ (n - r - 2)(n - r - 2 - 1) \ldots (n - r - 2 - (t - 2 - s)) ] \) is the number of ways we can permute the remaining numbers that lie outside \( b \) and \( c \).
12. \((n - 2 - t + 1)!\) is the number of ways we can permute the remaining numbers after taking out \(b\) and \(c\) and the numbers that lies within this interval.

13. \(\binom{n-r-t-1+s}{k-1}\) is the number of ways to choose either \(b\) or \(c\), and \(k-1\) other legal numbers, so as to generate a type A duplicate.

We will see that

1. \(\sum_{1}\) yields good values (i.e. smaller than \(\binom{n}{k}\)) for small examples, but there are exceptions.

2. \(\sum_{1}\) is almost exact for \(k = n - 1\).

3. \(\sum_{1}\) may cause double counting thus causing the number distinct subsequences to be underestimated.

Considering small values of \(n\), (say \(n = 4\) and \(k = 1, 2, 3\)) we see on the next page the comparison of the formula in Theorem 4.1 with the value of \(\binom{n}{k}\) in the Table 17. We can observe from the first row of the Table 17 that our formula gave us a value greater than \(\binom{n}{k}\) when \(k\) was 1. However, when \(k\) was 2, our formula gave us a smaller value than \(\binom{n}{k}\). We see that we do not always produce a number that is smaller than \(\binom{n}{k}\). We therefore believe that further research work needs to be carried out so that we can determine the true nature of the above Theorem.
Table 17: Comparison of the Main theorem with \(_{\binom{n}{k}}\).

<table>
<thead>
<tr>
<th>(k)</th>
<th>Formula</th>
<th>(_{\binom{n}{k}})</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>5.224</td>
<td>4</td>
</tr>
<tr>
<td>2</td>
<td>4.833</td>
<td>6</td>
</tr>
<tr>
<td>3</td>
<td>1.5</td>
<td>4</td>
</tr>
</tbody>
</table>

Derivation of the number of distinct subsequences: For example, taken

\[\pi = 41325, \quad k = 3,\]

we obtain from Table 18:

Table 18: Distinct subsequences in a permutation.

<table>
<thead>
<tr>
<th>312</th>
<th>213</th>
<th>123</th>
<th>132</th>
<th>321</th>
</tr>
</thead>
<tbody>
<tr>
<td>413</td>
<td>415</td>
<td>135</td>
<td>132</td>
<td>432</td>
</tr>
<tr>
<td>412</td>
<td>435</td>
<td>125</td>
<td>null</td>
<td>null</td>
</tr>
<tr>
<td>null</td>
<td>425</td>
<td>null</td>
<td>null</td>
<td>null</td>
</tr>
<tr>
<td>null</td>
<td>325</td>
<td>null</td>
<td>null</td>
<td>null</td>
</tr>
</tbody>
</table>

From table 18, given \(2^n - \sum_1\) (where \(\sum_1\) is repeats of Type A), we have the total left to be 2. In reality, there are 5 distinct subsequences of length 3 which are under the permutation 312, 213, 123, 132 and 321. However, after removing pairs and taking the type A duplicates out we are left with only 2 distinct categories (132 and 321). Because of double counting the total number of duplicates is 8 (these are under categories 312, 213 and 123). The established upper bound might not be a good upper bound because we might have a duplication caused by \(a\) and \(c\) and the same duplication caused by \(b\) and \(c\).
4.2 Special Cases of Theorem 4

4.2.1 Case 1, Successions with \( k = n - 1 \)

When \( k = n - 1 \) we can only have successions. No other values of \( r, t \) are possible, e.g. \( t \geq 2 \). We can only have \( t = 1, r = 0 \). The formula reduces to

\[
\frac{2(n - 1)(n - 1)(n - 2)!}{n!} \binom{n - 2}{n - 2} = \frac{2(n - 1)(n - 1)(n - 2)!}{n(n - 1)(n - 2)!} = \frac{2(n - 1)}{n} \approx 2 \\
\leq 2 \binom{n}{k} = \binom{n}{n - 1} = \binom{n}{1} = n. 
\]

(15)

This gives the expected number of successions to be 2, as per the result in Allison et al (2013). The formula is thus exact as stated earlier.

4.2.2 Case 1 - Successions with \( r = 0 \)

By way of a motivating example, assuming 5 and 8 are in the positions 15 and 16, if we pick 5, then we will not consider the numbers 6,7,8. In this case, we will be left with 16 choose 11 possibilities for the other numbers.

We define \( t = |c - b| = 1, (n - 1) = \text{number of starting positions and } 2(n - 1)(n - 2)!/n! = \text{number of ways of arranging the remaining numbers.} \) The formula thus yields
\[
\frac{2(n-1)(n-1)(n-2)!}{n!} \binom{n-2}{k-1} = \frac{2(n-1)(n-1)(n-2)!}{n(n-1)(n-2)!} \binom{n-2}{k-1} = \frac{2(n-1)}{n} \binom{n-2}{k-1} = \frac{2(n-1)}{n} (n-2)! = 2 \frac{(n-2)!}{(k-1)!(n-2-(k-1))!} = \frac{2n(n-1)(n-2)!k(n-k)}{n^2k(k-1)!(n-k-1)!(n-k)} = \frac{2n!k(n-k)}{n^2k!(n-k)!} = 2 \left( \frac{k}{n} \right) \left( \frac{n-k}{n} \right) \frac{n!}{(n-k)!k!} = 2 \left( \frac{k}{n} \right) \left( \frac{n-k}{n} \right) \frac{n}{k}. \tag{16}
\]
4.2.3 Case 2 - Near Successions with $r = 0$

First consider $t = 2, r = 0$, where the two numbers are next to each other but differ by 2. The formula gives

$$2 \frac{(n-1)(n-2)(n-2)!}{n!} \frac{(n-3)}{(k-1)}$$

$$= 2 \frac{(n-2)}{n} \frac{(n-3)}{(k-1)}$$

$$= 2 \frac{(n-2)}{n} \frac{(n-3)}{(k-1)}$$

$$= 2 \frac{(n-2)}{n} \frac{k}{n} \frac{(n-3)!n(n-1)(n-k)(n-k-1)}{k(k-1)!n(n-1)(n-k-2)(n-k)(n-k-1)}$$

$$= 2kn!(n-k)(n-k-1)$$

$$= 2k \frac{n-k(n-k-1)}{n} \frac{n!}{(n-k)!k!}$$

$$= 2k \frac{n-k(n-k-1)}{n} \frac{n}{n-1} \frac{(n-k)!k!}{n!}$$

$$= \frac{2k}{n} \left( \frac{n-k}{n} \right)^2 \left( \frac{n}{k} \right). \tag{17}$$

Now for $t = 3$ and $r = 0$, (for example 5,8). These numbers are next to each other but they differ by 3.

$$2 \frac{(n-1)(n-3)(n-2)!}{n!} \frac{(n-4)}{(k-1)}$$

$$= 2 \frac{(n-3)}{n} \frac{(n-4)}{(k-1)}$$

$$= 2 \frac{(n-3)}{n} \frac{(n-4)!n(n-2)(n-1)(n-k-2)(n-k-1)(n-k)(n-k-1)}{n.k(k-1)!(n-k-3)!n(n-2)(n-1)(n-k-1)(n-k-1)(n-k)(n-k-2)(n-k)(n-1)n}$$

$$= 2k.n!(n-k-2)(n-k-1)(n-k)$$

$$= 2k \frac{n-k}{n} \frac{n-k-1}{n-1} \frac{n-k-2}{n-2} \frac{(n-k)!k!}{n!}$$

$$\leq 2 \left( \frac{k}{n} \right) \left( \frac{n-k}{n} \right)^3 \left( \frac{n}{k} \right). \tag{18}$$
In general, the total contribution of the \( r = 0 \) case can be bounded by

\[
\leq 2 \left( \frac{k}{n} \right) \left(1 - \frac{k}{n}\right) \frac{1}{n-k} \binom{n}{k},
\]

yielding

**Proposition 4.2.** The total contribution of duplicates by adjacent terms is \( \leq 2 \left(1 - \frac{k}{n}\right) \binom{n}{k} \)

We combine these contributions for \( k \geq n/2 \) to yield

\[
2 \sum_{k=n/2+1}^{n} \left(1 - \frac{k}{n}\right) \binom{n}{k} = 2 \sum_{k=n/2}^{n} \binom{n}{k} - 2 \sum_{k=n/2+1}^{n} \frac{k}{n} \binom{n}{k}
\]

\[
= 2 \left(\frac{1}{2^2} \right) 2^n - 2 \sum_{k=n/2+1}^{n} \frac{k}{n} \frac{n!}{(n-k)!k!}
\]

\[
= 2^n - 2 \sum_{k=n/2+1}^{n} \frac{n-1}{k-1} = 2^n - 2 \sum_{j=n/2}^{n-1} \binom{n-1}{j}
\]

\[
= 2^{n-1}.
\]

(20)

4.2.4 Other Cases: Non-successions with \( r = 1 \)

In this next step, we consider the case \( r = 1 \) and \( t = 1 \). Thus the position difference for the two numbers \( b, c \) is 1 and the numerical difference is 1. An example
of such numbers are $5 \times 6$. The formula yields

$$\frac{2(n-1)(n-2)(n-2)!}{n!} \binom{n-3}{k-1}$$

$$= \frac{2(n-2)}{n} \binom{n-3}{k-1}$$

$$= \frac{2(n-2)(n-3)!}{n \cdot k(k-1)!} \frac{(n-1)n(n-k-1)(n-k)k}{(n-2-k)!n(n-1)(n-k-1)(n-k)}$$

$$= \frac{2}{n} \frac{n}{k} \frac{k(n-k-1)(n-k)}{n(n-1)}$$

$$\leq \frac{2}{n} \frac{n}{k} \frac{k(n-k)(n-k)}{n}$$.

$$\text{(21)}$$

Now we consider $r = 1, t = 2$. As an example, we consider the numbers $5 \times 7$, and note we can either have the number 6 between them or outside. When 6 lies outside of these two numbers, we will have another number to lie between this interval. This gives us two conditions (the case number in-value (eg 6) lies between them and the case when another number lies between them but we have to rule both the 6 and that number out since they can cause a repetition. This gives, via Theorem 4:

$$\frac{2(n-2)(n-2)}{n!} \left[ (n-3)! \frac{(n-3)}{k-1} + ((n-2)! - (n-3)!) \frac{(n-4)}{k-1} \right]$$

$$= \frac{2(n-2)(n-2)}{n!} \left[ (n-3)! \frac{(n-3)}{k-1} + (n-2)! \frac{(n-4)}{k-1} - (n-3)! \frac{(n-4)}{k-1} \right]$$

$$= \frac{2(n-2)(n-2)}{n!} \left[ (n-2)! \frac{(n-4)}{k-1} + (n-3)! \left( \frac{(n-3)}{k-1} - \frac{(n-4)}{k-1} \right) \right]$$

$$= \frac{2(n-2)(n-2)}{n!} \left[ (n-2)! \frac{(n-4)}{k-1} + (n-3)! \frac{(n-4)}{k-2} \right]$$

$$= \frac{2(n-2)(n-2)(n-2)!}{n!} \frac{(n-4)}{k-1} + \frac{2(n-2)(n-2)(n-3)!}{n!} \frac{(n-4)}{k-2}$$

$$\text{(22)}$$
For the first part, we have
\[
\frac{2(n-2)(n-2)(n-2)!}{n(n-1)(n-2)!} \binom{n-4}{k-1} = \frac{2(n-2)(n-2)(n-4)!}{n(n-1)(n-k-3)!(k-1)!}
\]
\[
= \frac{2(n-2)(n-2)(n-4)!}{n(n-1)(n-k-3)!(k-1)!}
\]
\[
\times \frac{n(n-1)(n-3)(n-k-2)(n-k-1)(n-k)}{n(n-1)(n-k-3)!(k-1)!k}
\]
\[
= \frac{2n!(n-2)(n-k-2)(n-k-1)k(n-k)}{k!n(n-1)(n-k)!(n-3)}
\]
\[
= 2\binom{n}{k} \frac{n-k}{n} \frac{n-k-1}{n-1} \frac{(n-k-2)(n-2)}{k} \frac{k}{(n-1)(n-3)} \frac{1}{n}
\]
\[
(23)
\]

Since \((n-k-2)(n-2) \leq (n-k-1)(n-3)\), (23) reduces to
\[
\frac{k}{n} \left( \frac{n-k}{n} \right)^2 \binom{n}{k}.
\]

For the 2nd part, note that
\[
\frac{2(n-2)(n-2)(n-3)!}{n!} \binom{n-4}{k-2} = \frac{2(n-2)(n-2)(n-4)!}{n!} \binom{n-4}{k-2}
\]
\[
= \frac{2(n-2)(n-2)(n-4)!}{n!} \binom{n-4}{k-2}
\]
\[
= \frac{2(n-2)(n-2)(n-3)!n(n-1)(n-3)(n-k-2)(n-k-1)(n-k)}{n(n-1)(n-k-1)!(n-k)!}
\]
\[
= \frac{k}{n} \frac{k-1}{n-1} \frac{n-k}{n} \frac{n-k-1}{n-1} \frac{1}{n-3} \binom{n}{k}
\]
\[
\leq 2 \left( \frac{k}{n} \right)^2 \left( \frac{n-k}{n} \right)^2 \frac{1}{n-3} \binom{n}{k}.
\]
\[
(24)
\]
5 CONCLUSION

In conclusion, we see that many of the earlier terms in the formula of Theorem 4 can be analyzed so as to be smaller that $\binom{n}{k}$. In future work, we need to complete this analysis and make further progress towards proving our conjecture.
BIBLIOGRAPHY


APPENDIX

\[ n = 3 \]

\begin{verbatim}
> result$count <- number
> result
perm  elements  count
 1  123  1, 12, 123  3
 2  132  1, 12, 21, 132  4
 3  213  1, 21, 12, 213  4
 4  312  1, 21, 12, 231  4
 5  231  1, 12, 21, 312  4
 6  321  1, 21, 321  3

> R=}((result$count)+1);# Number of Embedded sequences for each possible sequence of n
[1] 4 5 5 5 4
> table(R)# Yields the frequency table of the number of Embedded sequences for each possible sequence of n
R
4 5
2 4

> S=sum(R);# Total number of Embedded Sequences
[1] 28
> A=length(R);A
[1] 6
> prob.dist<-prop.table(table(R))
> prob.dist
R
4 5
0.2857143 0.7142857
> E<-(S/A);E
[1] 4.666667
> v=sum(unique((result$count)+1)^2*prob.dist)-E^2;v
[1] 0.2222222
\end{verbatim}
\[ n = 4 \]

```
> result
  perm elements count
1  1234  1, 12, 13, 23, 24  4
2  1243  1, 12, 21, 13, 23, 34  6
3  1324  1, 12, 21, 32, 13, 24  7
4  1423  1, 12, 21, 32, 13, 42  7
5  1342  1, 12, 21, 32, 14, 32  7
6  1432  1, 12, 21, 34, 13, 24  7
7  2134  1, 21, 12, 31, 13, 24  7
8  2143  1, 21, 12, 31, 13, 42  7
9  3124  1, 21, 12, 23, 13, 24  7
10 4123  1, 21, 12, 32, 13, 24  7
11 3142  1, 21, 12, 32, 13, 42  7
12 4132  1, 21, 12, 32, 13, 42  7
13 2314  1, 12, 21, 32, 12, 23  7
14 2413  1, 12, 21, 32, 13, 24  7
15 3214  1, 21, 12, 32, 13, 24  7
16 4213  1, 21, 12, 32, 13, 42  7
17 3412  1, 12, 21, 32, 12, 23  7
18 4312  1, 12, 21, 32, 13, 24  7
19 2341  1, 12, 21, 32, 12, 43  7
20 2431  1, 12, 21, 32, 13, 42  7
21 3241  1, 21, 12, 32, 12, 31  7
22 4231  1, 21, 12, 32, 13, 42  7
23 3421  1, 12, 21, 32, 12, 43  7
24 4321  1, 12, 21, 32, 13, 42  7
```

> R<-table(result$count) R# Number of Embedded Sequences for each possible sequence of n

> table (R ) # yields the frequency table of the number of Embedded sequences for each possible sequence of n

> s=sum(R) ; s # Total number of Embedded sequences

> A=length(R) ; A

> prob.dist<prop.table(table(R))

> prob.dist

> > E=-(s/A)

> E

> > v<-sum(unique(result$count)*1^2*prob.dist)-E^2

> v

> [1] 0.9097222
```
\( n = 5 \)

<table>
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<tr>
<th>perm</th>
<th>elements</th>
<th>count</th>
</tr>
</thead>
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<tr>
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<td>8</td>
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<tr>
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</tr>
<tr>
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<td>14</td>
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<td>14532</td>
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<td>11</td>
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<td>14532</td>
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<td>11</td>
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<tr>
<td>15432</td>
<td>1, 12, 21, 123, 132, 1232, 1231, 1423, 1342, 1423, 12345</td>
<td>11</td>
</tr>
<tr>
<td>12345</td>
<td>1, 12, 21, 123, 132, 1232, 1231, 1423, 1342, 1423, 12345</td>
<td>8</td>
</tr>
</tbody>
</table>

54
R=$((result$count)+1);# Number of Embedded Sequences for each possible sequence of n
12 15 15 14
] [40] 14 14 13 14 14 16 14 14 11 11 12 15 14 25 14 10 10 13 12 14 12 12 14 12 13 10 10 14 15 14 15 12 11 11 14
14 16 14 14
] [79] 13 14 14 14 15 15 12 14 9 11 12 11 10 12 12 11 11 9 9 12 14 14 13 11 12 12 14 12 15 13 13 15 11 13 11
11 10 11 11
] [118] 11 9 6
> table(R);# Yields the frequency table of the number of Embedded Sequences for each possible sequence of n
R
6 9 10 11 12 13 14 15 16
8 8 26 22 12 28 12 2
>
> S=sum(R);# Total number of embedded Sequences
[1] 1474
> A=length(R);
[1] 120
> prob.dist<-prop.table(table(R))
> prob.dist
R
6 9 10 11 12 13 14 15 16
0.01866667 0.06666667 0.06666667 0.21666667 0.18333333 0.10000000 0.23333333 0.10000000 0.01866667
>
> E<-S/A
> E
[1] 12.28333
>
> V<-(sum(unique((result$count)+1))2*prob.dist-E^2)
> V
[1] 4.196944
\[ n = 6 \]

```r
> table(R) # yields the frequency table of the number of Embedded Sequences for each possible sequence of n
R
7  11 13 14 15 16 17 18 19 20 21 22 23 24 25 26 27 28 29
2  8 14 20 18 18 38 48 52 72 70 36 80 76 66 6 44 24 8
>
> S=sum(R); S # Total number of Embedded Sequences
[1] 15224
> A=length(R); A
[1] 720
> prob.dist<-prop.table(table(R))
> prob.dist
R
7 0.002777778 0.011111111 0.019444444 0.027777778 0.025000000 0.025000000 0.080555556 0.066666667 0.072222222 0.100000000
11 0.097222222 0.050000000 0.111111111 0.105555556 0.091666667 0.083333333 0.061111111 0.033333333 0.011111111
13 0.097222222 0.050000000 0.111111111 0.105555556 0.091666667 0.083333333 0.061111111 0.033333333 0.011111111
14 0.097222222 0.050000000 0.111111111 0.105555556 0.091666667 0.083333333 0.061111111 0.033333333 0.011111111
15 0.097222222 0.050000000 0.111111111 0.105555556 0.091666667 0.083333333 0.061111111 0.033333333 0.011111111
16 0.097222222 0.050000000 0.111111111 0.105555556 0.091666667 0.083333333 0.061111111 0.033333333 0.011111111
17 0.097222222 0.050000000 0.111111111 0.105555556 0.091666667 0.083333333 0.061111111 0.033333333 0.011111111
18 0.097222222 0.050000000 0.111111111 0.105555556 0.091666667 0.083333333 0.061111111 0.033333333 0.011111111
19
>
> E<-S/A
> E
[1] 21.14444
>
> V<-sum(unique(result$count)+1)^2*prob.dist-E^2
> V
[1] 25.63469
> ```
\[ n = 7 \]

\[ s = \text{sum}(R); s \quad \text{# Total number of Embedded Sequences} \]
\[ A = \text{length}(R); A \]
\[ n = \text{prob.dist}\left(\text{prop.table}(table(R))\right) \]
\[ n = 8 \]

\[ R \]
\[ 8 \quad 13 \quad 16 \quad 17 \quad 19 \quad 20 \quad 21 \quad 22 \]
\[ 23 \quad 0.0003968254 \quad 0.0015873016 \quad 0.0023809524 \quad 0.0055355556 \quad 0.0051587302 \quad 0.0019841270 \quad 0.0015873016 \quad 0.0130952381 \quad 0.0150793651 \]
\[ 24 \quad 0.0142857143 \quad 0.0103174603 \quad 0.0202380952 \quad 0.0174803175 \quad 0.0206394266 \quad 0.0333333333 \quad 0.0388888889 \quad 0.0285714286 \quad 0.0373015873 \]
\[ 25 \quad 0.0365079365 \quad 0.0380952381 \quad 0.0507936508 \quad 0.0412698413 \quad 0.0575396825 \quad 0.0320315873 \quad 0.0151873016 \quad 0.0428571429 \quad 0.0333333333 \]
\[ 26 \quad 42 \quad 43 \quad 44 \quad 45 \quad 46 \quad 48 \quad 49 \]
\[ 50 \quad 0.0403174603 \quad 0.0703015873 \quad 0.0301587302 \quad 0.0800396825 \quad 0.0202380952 \quad 0.0230158730 \quad 0.0230158730 \quad 0.0230158730 \quad 0.0230158730 \quad 0.0079365079 \]
\[ 51 \quad 52 \quad 53 \quad 54 \quad 55 \quad 56 \]
\[ 0.0261904762 \quad 0.0142857143 \quad 0.0703015873 \quad 0.0047019048 \quad 0.0047019048 \quad 0.0019841270 \]
\[ E = (S/A); E \]
\[ 37.60397 \]
\[ V = \text{sum}(\text{unique}(\text{result$count})+1)^2*\text{prob.dist})-E^2; V \]
\[ 76.07927 \]

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